



# Comparing latent inequality with ordinal health data



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## Motivation

### Deaton and Paxson (1998a)

Empirical questions

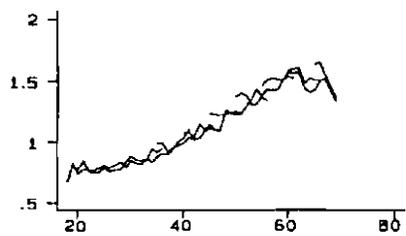
1. Health better? (inequality “across socioeconomic groups”)
2. Health more dispersed? (“whether inequality in health  $\uparrow$  with age”)

- Ordinal SRHS (self-reported health status): “excellent,” “good,” etc.
- SRHS benefits: 1) “useful over the complete adult life cycle,” 2) strongly correlated with objective measures, 3) widely available (PSID, NHIS, etc.), 4) synthesizes all health dimensions.

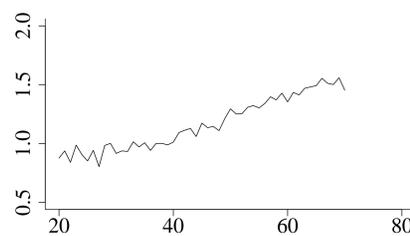
Deaton and Paxson (1998a,b): SRHS “variance”  $\uparrow$  with age.

But: same pattern in simulation of pure location shift of latent health.

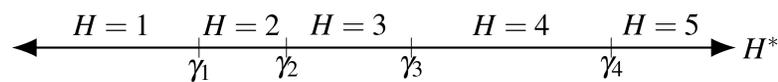
Var(SRHS) in Figure 10.4 (1998b)



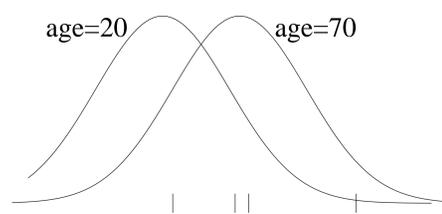
Var(SRHS) in Simulation



Latent health  $H^*$ , ordinal  $H$ , thresholds  $\gamma_j$ :



Simulation DGP: for ages  $a = 20, \dots, 70$ , sample 1000 iid  $N(\mu_a, 1)$  each for increasing  $\mu_a$ , convert to ordinal using fixed thresholds.



Problems with current methods in literature

1. “Cardinal sin” (treating excellent=1, ..., poor=5).
2. Treating  $H^*$  as discrete.
3. Unrealistic parametric models. (See also Bond and Lang, 2018)

- SRHS inequality: mostly Problem 2, including Allison and Foster (2004) “median-preserving spread.” Inequality indexes further rely on choice of weights, mostly ignore inference.
- Partial identification: Stoye (2010) bounds spread parameters given CDF bounds.

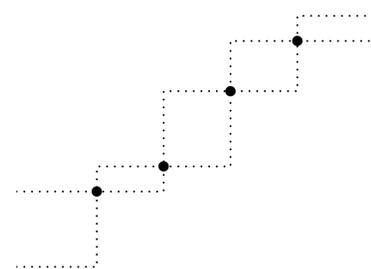
## Identification

- “Restricted SD1”: Atkinson (1987).
- Location-scale model  $\implies$  single (or no) CDF crossing.
- Thresholds shift by: zero? constant? idiosyncratic? (Lindeboom and van Doorslaer, 2004; Hernández-Quevedo, Jones, and Rice, 2005).

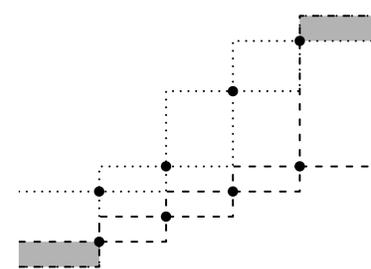
Proposition 2 [SD1]: assuming same thresholds,

- (i) Latent SD1  $\implies$  ordinal SD1 (testable implication).
- (ii) Location-scale + ordinal SD1  $\implies$  latent “restricted SD1.”
- (iii) Ordinal “super-SD1”  $\implies$  latent restricted SD1.

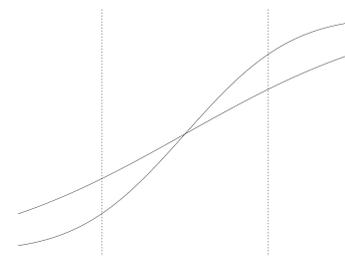
CDF bounds



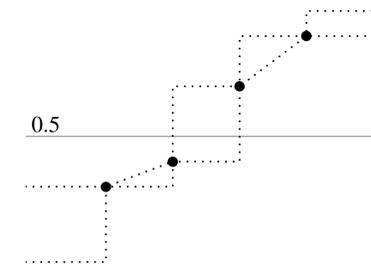
Proposition 2(iii)



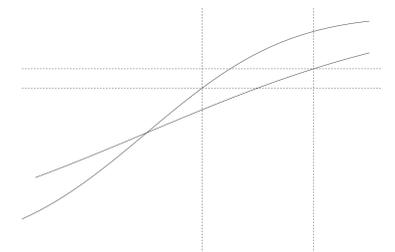
Proposition 3(i) [IQR $\uparrow$ ];  $\gamma_j$  may shift by constant



CDF bounds if unimodal/symmetric



Proposition 4(i,ii) [IQR $\uparrow$ ]



Location-scale  $\implies$  extrapolate single IQR to overall “scale.”

## Inference

- Statistical inference on ordinal relationships in Propositions 2–4.
- Ordinal  $H_2 \text{ SD}_1 H_1 \iff \mathbb{E}[\mathbb{1}\{H_2 \leq j\} - \mathbb{1}\{H_1 \leq j\}] \leq 0, j = 1, 2, 3, 4$ . Can use moment inequality tests (Andrews and Soares, 2010, etc.)
- Other relationships: more complicated (unions of intersections of ...).
- Bayesian: Dirichlet-multinomial model. Posteriors for all relationships computed easily, simultaneously, and coherently (e.g., probabilities of SD1 and non-SD1 sum to 1).

## References

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